# Reliability Estimate of Technical Systems Containingcomposite Materials through Analysis of the Concurrent Risks

## Nikolay Iv. Petrov<sup>1</sup>, Dinara Dzh. Baskanbayeva<sup>2</sup>

Technical University-Sofia, The Kazakh National Technical Research University after K.I.Satpaev (KazNTRU)- Republic of Kazakhstan, Almaty city

Abstract: The technical systems containing composite materials are risk te-chnical systems (RTS). The investigation of use reliability of RTS is correlation with analyze of the exist mutual concurrent risks. This analyze demand collecting of the statistical information (sample of data) for this conduct. In row cases howe-ver this sample is determinated censorship (unfully). The general situation with random censorship (unfully) extract is characterized for RTS. This is provoked by fact that observed at technical usage article, could be a failure by two or more independence reason. Whit other words, because of article have take  $\Gamma$  independence motive provoke failures. (concurrent risks), but is display one of them. At this the presence of aunfully sample could be get, while same at different kind failures.

Therefore, observed time a failure of article by j -th kind (j = 1, ..., r) is determi-nated by realization of convention random quantity (by condition, at isn't step a fai-lure at other kind).

Keywords: reliability estimate; technical systems; composite materials

## 1. Introduction

In this study it is offered an estimate of the probability, article serviceability in moment  $t_i$ , to failure in interval of time  $(t_i, t_{i+1})$  by i- th cause under action of the catastrophic functional failure provoking by a risks. In reference [6] this probability is called gross probability. Let in first moment of time  $t_0$  under observer is find technical systems from  $n_0$  on regeneration article has probability for reliability work  $P_{0i}$ . For i- th time interval  $(t_{i-1}, t_i)$  with continuation  $\Delta t = t_i - t_{i-1}$ , every a article, non failureearlier, can be faultless by j-th reason with probability reliability work (PRW)  $P_{\text{EPi}}(\Delta t)$  de-terminate by equation:

$$P_{RWi}\left(\Delta t\right) = 1 - \sum_{j=1}^{r} Q_{ij}\left(\Delta t\right). \quad (1)$$

### 2. Presentation

In every interval  $\Delta t$  advance  $d_{ij}$  failures with general case  $d_i = \sum_{j=1}^r d_{ij}$ , like as that number of articles  $n_{j+1}$ ,

remains serviceability in start of the (i+1) - th interval is

$$n_{i+1} = n_0 - \sum_{j=1}^{i} d_j$$
 or  $n_i = \sum_{j=1}^{r} d_{ij} + n_{i+1}$ 

Consequently, at  $n_i > 0$  articles, serviceability by the moment  $t_i$  is characterized: "Conditional distribution with the number of failures in interval  $\Delta t$  and number of articles  $n_{i+1}$ , serviceability in the end of interval, is polynomiality" [1].

The distribution of probability is according[4]:

$$P\{d_{ij}, n_{i+1}|n_i\} = \left\lfloor n_i ! / \left(n_{i+1} ! . \prod_{j=1}^r d_{ij}\right) \right\rfloor \cdot P_i^{n_i+1} \cdot \prod_{i=1}^r Q_{ij}^{d_{ij}}.$$
(2)

The estimate of probability for failure  $\hat{Q}_{ij}(\Delta t)$  and PRW

 $\hat{P}_{EPi}(\Delta t)$ , are de-terminated by the results for observer of articles working in time intervals W of number (W is determinate by condition  $n_W > 0$  and  $n_{W+1} \ge 0$ ), getting with *functional of reality like* L has been case:

$$L = \prod_{i=1}^{W} P\{d_{ij}, n_{i+1} | n_i\}.$$
 (3)

The estimate of this value is:

$$\hat{P}_{_{BPi}}(\Delta t) = n_{_{i+1}}/n_i; \ \hat{Q}_{_{ij}}(\Delta t) = d_{_{ij}}/n_i.$$
 (4)

The mathematical expectation of the value  $\hat{Q}_{ij}(\Delta t)$  and  $\hat{P}_{RW,i}(\Delta t)$ , could be obtained with calculation:

# Volume 6 Issue 10, October 2018 <u>www.ijser.in</u>

# Licensed Under Creative Commons Attribution CC BY

Index Copernicus Value (2015): 56.67 | Impact Factor (2017): 5.156

$$M\left(\hat{Q}_{ij}\right) = M\left(d_{ij}/n_{i}\right) = M\left(\frac{1}{n_{i}} \cdot M\left(d_{ij}|n_{i}\right)\right) - M\left(\frac{1}{n_{i}} \cdot Q_{ij}.n_{i}\right) = Q_{ij};$$
$$M\left(\hat{P}_{RWi}\right) = P_{RWi}.(5)$$

By conditional calculations, it could be obtained the dispersion of the value:

$$M\left(\hat{Q}_{ij}^{2}\right) = M\left(d_{ij}^{2}/n_{i}^{2}\right) = M\left(\frac{1}{n_{i}^{2}} \cdot M\left(d_{ij}^{2}/n_{i}\right)\right) = M\left\{\frac{1}{n_{i}^{2}} \cdot \left[n_{i}^{2} \cdot Q_{ij}^{2} + n_{i} \cdot Q_{ij} \cdot (1 - Q_{ij})\right]\right\} = Q_{ij}^{2} + M\left(\frac{1}{n_{i}}\right) \cdot Q_{ij} \cdot (1 - Q_{ij}).$$
(6)

From equation (6) with (5) it is obtained the folloing new equation by the kind:

$$D(\hat{Q}_{ij}) = M\left(\frac{1}{n_i}\right) \cdot Q_{ij} \cdot \left(1 - Q_{ij}\right). \quad (7)$$

Like this it is determinated the dispersion for value of  $P_{RW}$ 

$$D(\hat{P}_{RWi}) = M\left(\frac{1}{n_i}\right) \cdot P_{RWi} \cdot (1 - P_{RWi}). \quad (8)$$

The expression  $M(1/n_i)$  in equations (7) and(8) insists on approximation. For doing this approximation it is obtained the following estimation. Let exa-mine the equation:

$$n_{i} = M(n_{i}) \{ 1 + [n_{i} - M(n_{i})] / M(n_{i}) \} = M(n_{i}) (1 + \Delta)$$
(9)

correctly with execution of a condition  $M(\Delta) = 0$ . From equality (9) follows after condition transformation new equation:

$$M\left(\frac{1}{n_i}\right) = \frac{1}{M(n_i)} \cdot M\left(\frac{1}{1+\Delta}\right).$$
(10)

Let use identity  $\frac{1}{1+\Delta} \equiv 1 - \Delta + \Delta^2 / (1+\Delta)$  [4], and

to put in equation (10). After transformation of the equation (10) it is obtained the inequality:

$$M\left(\frac{1}{n_{i}}\right) = \frac{1}{M\left(n_{i}\right)} \cdot \left\{1 + M\left[\Delta^{2}/1 + \Delta\right]\right\} \ge \frac{1}{M\left(n_{i}\right)}$$
(11)

in which random quantity  $\Delta^2$  and  $1+\Delta$  take only one negative and positive value.

The expand of the function  $1/n_i$  in Taylor's series in neighborhood of a point  $M(n_i)$  with accuracy by fort member and respective mathematical expectation, determinates the equation:

$$M\left(\frac{1}{n_{i}}\right) = \frac{1}{M(n_{i})} \cdot \left\{1 + D(n_{i})/M^{2}(n_{i}) + \upsilon \left[D(n_{i})/M^{2}(n_{i})\right]\right\}$$
(12)

where second member in brackets is relative dispersion of  $n_i$ , and third member is quantity of the higher ordinate.

The formulas for mathematical expectation and dispersion, have kind ac-cording [2, 3]:

$$M(n_{i}) = n_{0} \cdot \prod_{j=1}^{i-1} P_{RWj} = n_{0} \cdot P_{0i},$$
  

$$D(n_{i}) = n_{0} \cdot P_{0i} \cdot (1 - P_{0i}),$$
  
following new equation:

$$D(n_i)/M^2(n_i) = (1 - P_{0i})/(n_0 \cdot P_{0i}) \cdot (13)$$

With high values of  $n_0$ , equation (13) will be striving by a value zero. Be-cause of this, the equations (11) and(13) the mathematical dependence can be shown by the next expression:

$$\lim_{n\to\infty} M(1/n_i) = 1/M(n_i),$$

following, that the equation (7) and (8) could be approximated whit expressions:

$$D(\hat{Q}_{ij}) \approx (1/n_0.P_{0i}).Q_{ij}.(1-Q_{ij}).(14)$$
$$D(\hat{P}_{Rm}) \approx (1/n_0.P_{0i}).P_{i}.(1-P_i).(15)$$

For determination of the efficient of the estimate  $Q_{ij}$  it is necessary to finding the low limit of the their dispersion. By the inequality of the Rao-Kramer fol-lowing, that

$$D(\hat{Q}_{ij}) \geq -[M(\partial^2 \ln L/\partial Q_{ij}^2)]^{-1}.$$

Simultaneously by the equations (14) and (15) it is determinated after respective transformation equality:

$$-\left[M\left(\partial^2 \ln L/\partial Q_{ij}^2\right)\right] = n_0 \cdot P_{0i} \cdot \left(P_i + Q_{ij}\right)/P_i \cdot Q_{ij} \cdot (16)$$

#### Volume 6 Issue 10, October 2018

<u>www.ijser.in</u> Licensed Under Creative Commons Attribution CC BY By the equation (16) following, that *low limit*  $V_{LL}(\hat{Q}_{ij})$  of

an estimate of the dispersion is equal to:

$$V_{LL}\left(\hat{Q}_{ij}\right) = P_{RWi} \cdot Q_{ij} / \left[n_0 \cdot P_{0i} \cdot \left(P_{RWi} + Q_{ij}\right)\right] (17)$$

After confront of the equation (14) and (17) it is obtained as a final result, the equation:

$$V(\hat{Q}_{ij}) - V_{LL}(\hat{Q}_{ij}) = (Q_{ij}/n_0.P_{oi}).(1/(P_{RWi} + Q_{ij}) - 1) \ge 0$$
(18)

From the equation (18) follows, that the estimate  $\hat{Q}_{ij}$  is effective only in case, when it is possible the condition  $P_{RW,i} + Q_{ij} = 1$ , i.e. when actions only one risk. At the better values of  $n_0$  the left side of the equation (18) is striving to zero. Consequently, the estimate  $\hat{Q}_{ij}$  is the asymptotically effective.

## 3. Conclusion

- 1) In the present study it is determinate analysis for use reliability of the RTScontaining composite materials.
- 2) This analysis is correlation with estimation of the existence mutual con-current risks.
- 3) It is proven, that the estimate of the probability reliability

work  $\hat{P}_{RW,i}$  and probability of failure  $\hat{Q}_{ij}$  are only estimates, on removing by the real value.

# References

- Petrov, N.Reliability"s Investigations of Risk Technical Systems.Monograph, PH ,,J. Uchkov", Bulgaria, Third Additionally Publication, 2018, ISBN 978-954-391-120-2, pp. 78-82.
- [2] Гиндев, Е. Въведение в теорията и практиката на надеждността. Част 1. Основи на приложната надеждност. Акад. изд. "Проф. Марин Дри-нов", София, България, 2000.
- [3] Гиндев, Е. Въведение в теорията и практиката на надеждността. Част 2. Осигуряващи процедури в надеждността. Акад. изд. "Проф. Марин Дринов", София, България, 2002.
- [4] *Петров, Н.*Надеждностни изследвания на рискови технически системи. Монография, ИК "Учков", България, 2007, ISBN 978-954-9978-92-6, р. 90.

# **Author Profile**



**Nikolay Iv. Petrov,** Academician Professor, DScTech, DScEcon Nikolay Ivanov Petrov was born in the city of Yambol, Bulgaria in 1953. He has graduated from the National Military University "VasilLevski", faculty of Aviation in 1977 and from

the University of National and World Economy – Sofia in 2005. He has defended a total of three dissertations – one thesis for his PhD

Volume 6 Issue 10, October 2018 <u>www.ijser.in</u> Licensed Under Creative Commons Attribution CC BY

degree and two dissertations for the degrees of DScTech (2001) and DScEcon (2015). He is also a member of the Union of independent Bulgarian writers. He has published 25 monographs and studies, 25 textbooks and 5 literary works.



**Dinara Dzh. Baskanbayeva,** Doctoral student of the first course of the Kazakh National Technical University named after Satpayev, Baskanbayeva Dinara Jumabaevna was born in the city of Ucharal in 1979.In 2002 the year she graduated from Kazakh

"machinery and equipment of oil and gas fields". In 2013 graduated with a master degree, specialty "Technological machines and equipment". She is a leading expert on the technical science of grant funding at the University.Has 10 scientific articles. 2 of them are publishing in foreign countries and 1 indexed impact factor Scopus.